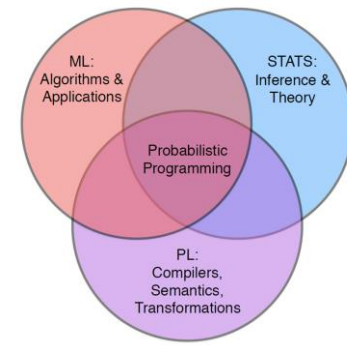


Probabilistic & **A**pproximate **C**omputing

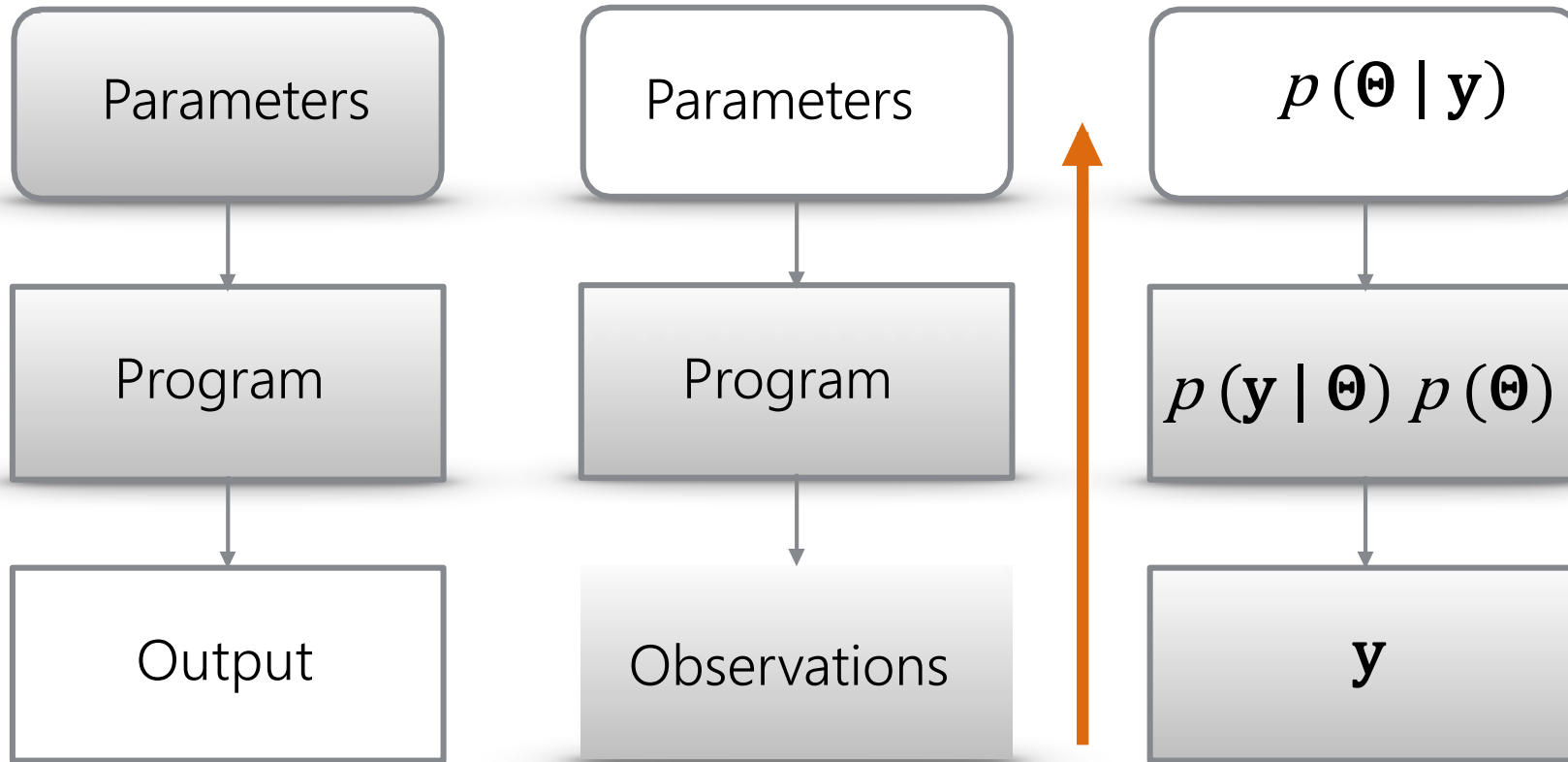
Sasa Misailovic

UIUC

Intuition



Inference



Programming

Probabilistic
Programming

Statistics

* Slide from Frank Wood

Example: Bernoulli Program

(

X	Y
--	--

, 1.0)

X := Bernoulli(0.7);

(

X	Y
True	--

, 0.7), (

X	Y
False	--

, 0.3)

Y := not X;

(

X	Y
True	False

, 0.7), (

X	Y
False	True

, 0.3)

Example: Bernoulli Program

(

X	Y
--	--

, 1.0)

X := Bernoulli(0.7);

(

X	Y
True	--

, 0.7), (

X	Y
False	--

, 0.3)

Y := Bernoulli(0.7);

(

X	Y
True	True

, 0.49), (

X	Y
True	False

, 0.21), (

X	Y
False	True

, 0.21), (

X	Y
False	False

, 0.09)

Example: Bernoulli Program

(

X	Y
--	--

, 1.0)

X := Bernoulli(0.7);

Y := Bernoulli(0.7);

(

X	Y
True	True

, 0.49), (

X	Y
True	False

, 0.21), (

X	Y
False	True

, 0.21), (

X	Y
False	False

, 0.09)

condition (X == True);

return Y;

(

X	Y
True	True

, 0.49/0.7), (

X	Y
True	False

, 0.21/0.7),

Example: Bernoulli Program

(

X	Y
--	--

, 1.0)

X := Bernoulli(0.7);

Y := Bernoulli(0.7);

(

X	Y
True	True

, 0.49), (

X	Y
True	False

, 0.21), (

X	Y
False	True

, 0.21), (

X	Y
False	False

, 0.09)

condition (X == Y);

return Y;

(

X	Y
True	True

, 0.49/0.58), (

X	Y
False	False

, 0.09/0.58)

Example: Bernoulli Program

(

X	Y
--	--

, 1.0)

X := Bernoulli(0.7);

Y := Bernoulli(0.7);

(

X	Y
True	True

, 0.49), (

X	Y
True	False

, 0.21), (

X	Y
False	True

, 0.21), (

X	Y
False	False

, 0.09)

factor (X ? 0 : -1);

$e^0 = 1$
 $e^{-1} = 0.37$

(

X	Y
True	True

, 0.49), (

X	Y
True	False

, 0.21), (

X	Y
False	True

, 0.077), (

X	Y
False	False

, 0.033)

Example: Bernoulli Program

factor \Rightarrow *condition*

(

X	Y
--	--

, 1.0)

X := Bernoulli(0.7);

Y := Bernoulli(0.7);

(

X	Y
True	True

, 0.49), (

X	Y
True	False

, 0.21), (

X	Y
False	True

, 0.21), (

X	Y
False	False

, 0.09)

factor (X ? 0 : $-\infty$);

$e^0 = 1$
 $e^{-\infty} = 0$

(

X	Y
True	True

, 0.49), (

X	Y
True	False

, 0.21), (

X	Y
False	True

, 0), (

X	Y
False	False

, 0)

Example: Bernoulli Program

(

X	Y
--	--

, 1.0)

X := Bernoulli(0.7);

Y := Bernoulli(0.7);

(

X	Y
True	True

, 0.49), (

X	Y
True	False

, 0.21), (

X	Y
False	True

, 0.21), (

X	Y
False	False

, 0.09)

condition (X == Y);

return Y;

(

X	Y
True	True

, 0.49/0.58), (

X	Y
False	False

, 0.09/0.58)

Example: Bernoulli Program

(

X	Y
--	--

, 1.0)

X := Bernoulli(0.7);

Y := Bernoulli(0.7);

(

X	Y
True	True

, 0.49), (

X	Y
True	False

, 0.21), (

X	Y
False	True

, 0.21), (

X	Y
False	False

, 0.09)

factor (X ? 0 : -1);

$e^0 = 1$
 $e^{-1} = 0.37$

(

X	Y
True	True

, 0.49), (

X	Y
True	False

, 0.21), (

X	Y
False	True

, 0.077), (

X	Y
False	False

, 0.033)

Example: Bernoulli Program

factor \Rightarrow *condition*

(

X	Y
--	--

, 1.0)

X := Bernoulli(0.7);

Y := Bernoulli(0.7);

(

X	Y
True	True

, 0.49), (

X	Y
True	False

, 0.21), (

X	Y
False	True

, 0.21), (

X	Y
False	False

, 0.09)

factor (X ? 0 : $-\infty$);

$e^0 = 1$
 $e^{-\infty} = 0$

(

X	Y
True	True

, 0.49), (

X	Y
True	False

, 0.21), (

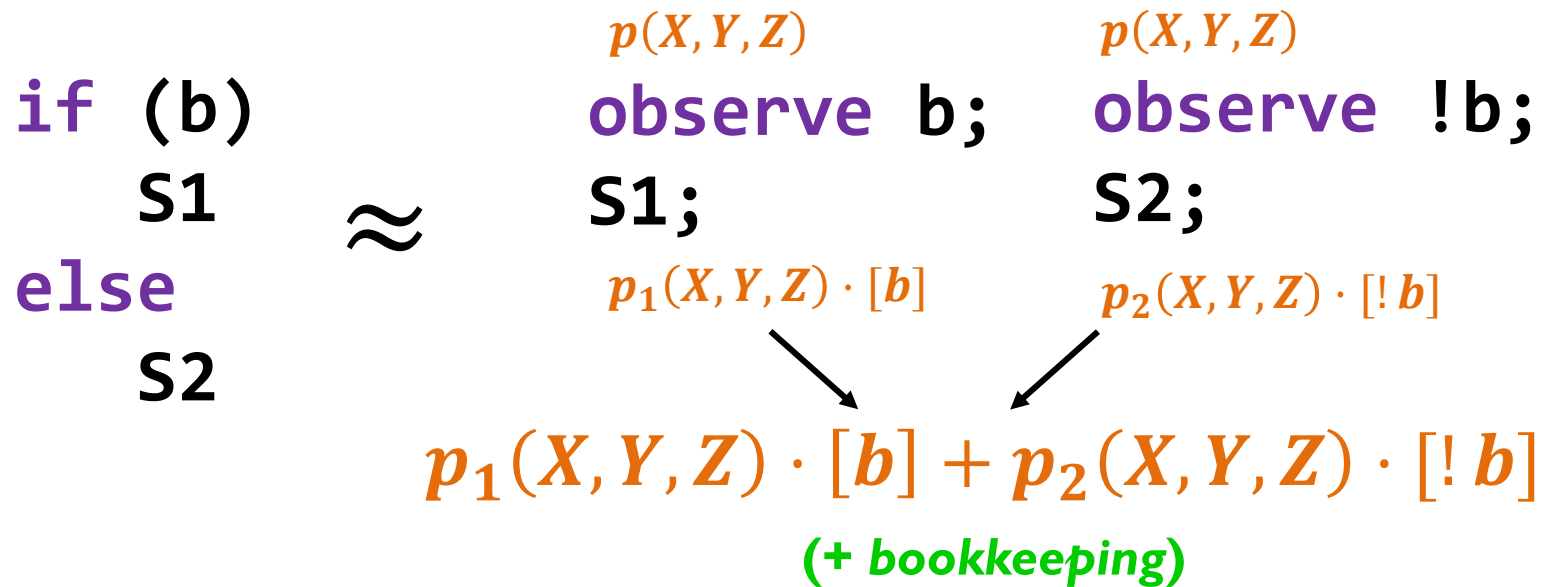
X	Y
False	True

, 0), (

X	Y
False	False

, 0)

Handling Conditionals



Handling Loops

```
N = Random(...)  
for i in [0..N)  ≈  
    S
```

```
p(X,Y,Z)  
N = Random(...)  
assert (N <= Nmax);  
for i in [0 .. Nmax) {  
    if (i < N)  
        S  
}
```

Probability of state: $p_{loop}(X, Y, Z) \cdot p_{N \leq N_{max}}$

Probability of error: $1 - p_{N \leq N_{max}}$

Unrolling factor depends on property to check

e.g. `assert (X > 0) withPr > 0.99`

Beyond Bayesian Net Models

Geometric Distribution: Probability of the number of Bernoulli trials to get one success

```
var geometric = function() {  
  return flip(.5) ? 0 : geometric() + 1;  
}
```

```
var dist = Infer({method: 'enumerate', maxExecutions: 10},  
                 geometric);  
viz.auto(dist);
```

Exact Inference

Naïve approach: Compute $P(x_1, x_2, \dots, x_n)$

Better approach:

Take advantage of (conditional) independencies

- Whenever we can expose conditional independence, e.g., $P_{X_1 X_2 | X_3}(x_1, x_2 | x_3) = P_{X_1 | X_3}(x_1 | x_3) \cdot P_{X_2 | X_3}(x_2 | x_3)$ the computation is more efficient
- Instead of the full table, keep smaller tables (for the individual variable distributions)

Compute distributions from parents to children

Complexity of Exact Inference

Number of variables: n

Naïve enumeration: complexity is $O(2^n)$

Variable Elimination: if the maximum number of parents of the nodes is $k \in \{1, \dots, n\}$, then the complexity is $n \cdot O(2^k)$.

For many models this is a good improvement, but always possible to construct pathological models.

Exact inference: Pragmatics

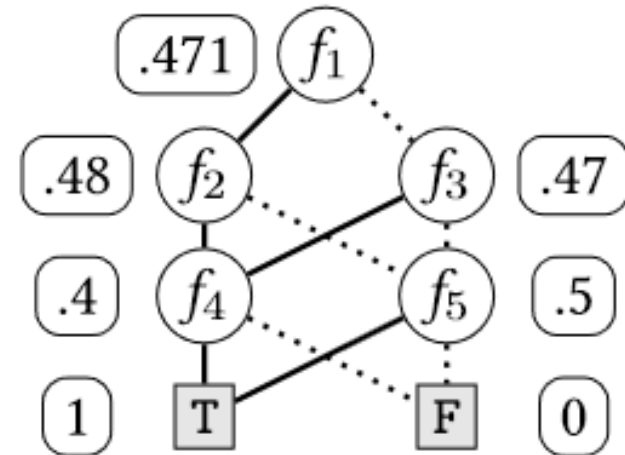
- Use dynamic programming where possible
- The algorithms for inference in Bayesian Networks are directly applicable if the program can be translated to one
- Leverage techniques from model checking
see *model counting*: compute the number of models (distinct variable assignments) for which some formula evaluates to true

Example of Model Checking

<http://dicelang.cs.ucla.edu>

```
1 let x = flip1 0.1 in
2 let y = if x then flip2 0.2 else
3   flip3 0.3 in
4 let z = if y then flip4 0.4 else
5   flip5 0.5 in z
```

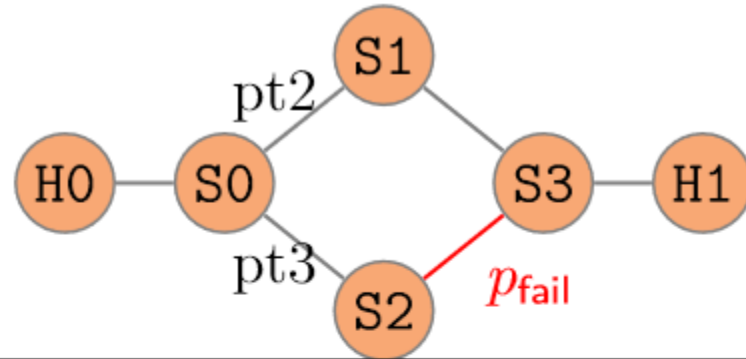
(a) Example Dice program.



(b) Compiled BDD with weighted model counts.

Fig. 1. Illustration of compiling a Dice program that exploits factorization.

Example: Network Reliability *Bayonet, PLDI 2018*

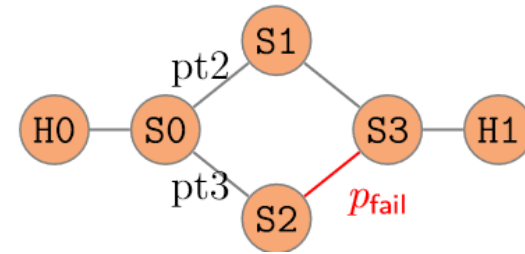


```
1 topology{
2   nodes{ H0 , H1 , S0 , S1 , S2 , S3 }
3   links{
4     (H0 ,pt1) <-> (S0 ,pt1) ,
5     (S0 ,pt2) <-> (S1 ,pt1) ,
6     (S0 ,pt3) <-> (S2 ,pt1) ,
7     (S1 ,pt2) <-> (S3 ,pt1) ,
8     (S2 ,pt2) <-> (S3 ,pt2) ,
9     (S3 ,pt3) <-> (H1 ,pt1)
10  }
11 }
```

To try out:

<https://dl.acm.org/doi/10.1145/3211997/full/>

Example: Network Reliability

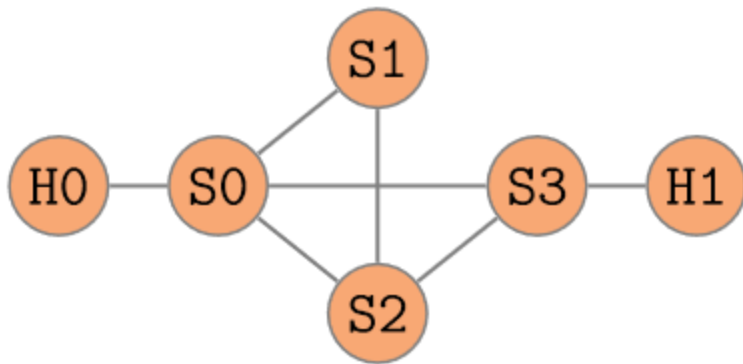


```
1 def h0(pkt, port){
2   fwd(1);
3 }
4 def h1(pkt, port)
5   state arrived(0)
6 {
7   arrived=1;
8   drop;
9 }
10 def s0(pkt, port){
11   if flip(1/2){
12     fwd(2);
13   }else{
14     fwd(3);
15   }
16 }
```

```
1 def s1(pkt, port){
2   fwd(2);
3 }
4 def s2(pkt, port)
5   state fail(
6     flip(1/1000))
7 {
8   if fail {
9     drop;
10  } else {
11    fwd(2);
12  }
13 }
14 def s3(pkt, port){
15   fwd(3);
16 }
```

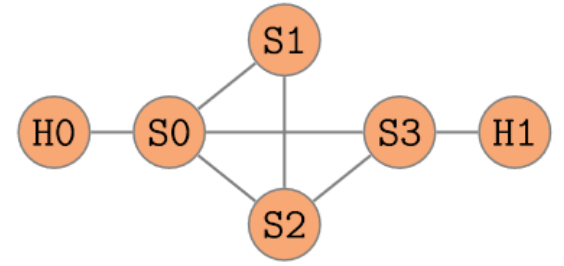
```
1 query probability(arrived@H1); // 0.9995
```

Example: Network Congestion



```
1 topology{
2   nodes{ H0 , H1 , S0 , S1 , S2 , S3
3     }
4   links{
5     (H0 , pt1) <-> (S0 , pt1) ,
6     (S0 , pt2) <-> (S1 , pt1) ,
7     (S0 , pt3) <-> (S3 , pt1) ,
8     (S0 , pt4) <-> (S2 , pt1) ,
9     (S1 , pt2) <-> (S2 , pt2) ,
10    (S2 , pt3) <-> (S3 , pt2) ,
11    (S3 , pt3) <-> (H1 , pt1)
12  }
13 queue_capacity 2;
```

Example: Network Congestion



```
1 def h0(pkt, port) state pkt_count(0){
2   if pkt_count < 3 {
3     new; pkt_count = pkt_count + 1; fwd(1);
4   } else { drop; }
5 }
6 def s0(pkt, port){
7   if flip(1/3){ fwd(1); }
8   else if flip(1/2){ fwd(2); }
9   else{ fwd(3); }
10 }
11 def s1(pkt, port){ fwd(2); }
12 def s2(pkt, port){ fwd(3); }
13 def s3(pkt, port){ fwd(3); }
14 def h1(pkt, port) state pkt_count(0){
15   pkt_count = pkt_count + 1; drop;
16 }
17 query probability(pkt_count@H1 < 3); // ?
```

Scheduler:

```
1 RunSw:=0, FwdQ:=1;
2
3 def scheduler(){
4   actions := []: (ℝ × ℝ) []; // list of actions
5   for i in [0..k) { // k: number of nodes
6     if (Q_in@i).size() > 0 {
7       actions ~= [(RunSw, i)]; } // run node
8     if (Q_out@i).size() > 0 {
9       actions ~= [(FwdQ, i)]; // forward packet
10    }
11  }
12  n := actions.length; // pick action u.a.r.:
13  return actions[uniformInt(0,n-1)];
14 }
```

// 2663191003150725233/5997013881313296384 ≈ 0.444

Continuous Models

The distributions in the program are continuous

- We are computing the log-likelihood of the trace

Doing 'hard' observations is ineffective: the probability of each observation is 0.

- Instead, use 'factor': the 'soft' version of conditioning

TrueSkill Model

Player Skill: Initially, we assume all players have a similar (randomly assigned) skill, centered around starting point

$$Skill_{player} \sim Gaussian(100, 10)$$

Player Performance: it is based primarily on skill, but in every game it can be higher or lower, based on the player's inspiration

$$Performance_{player} \sim Gaussian(Skill, 15)$$

Tournament Scores: Each player plays against others, we record the victory if the player's performance was higher than their opponent's

$$Performance_{player\#1} > Performance_{player\#2}$$

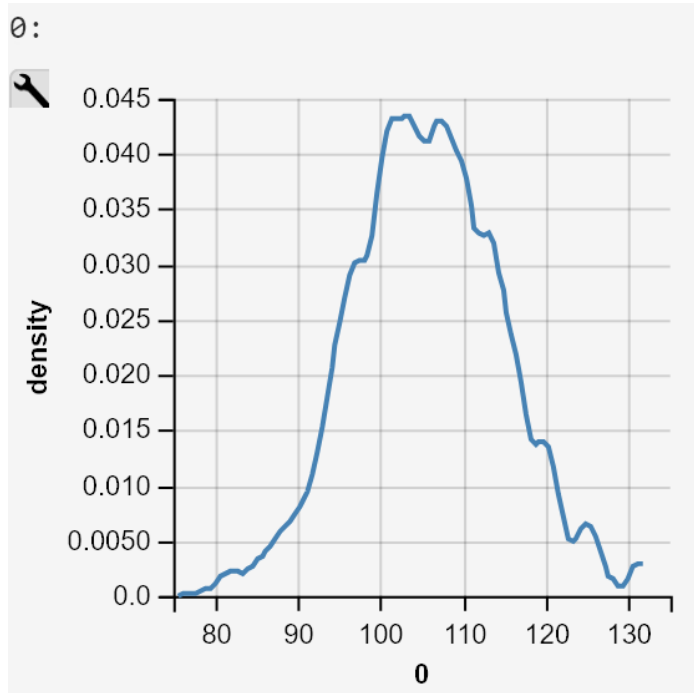
TrueSkill Example: 3 Players

```
var trueSkill = function(){  
  
    var skillA = gaussian(100, 10);  
    var skillB = gaussian(100, 10);  
    var skillC = gaussian(100, 10);  
  
    var perfA1 = gaussian(skillA, 15), perfB1 = gaussian(skillB, 15);  
    condition (perfA1 > perfB1);  
  
    var perfB2 = gaussian(skillB, 15), perfC2 = gaussian(skillC, 15);  
    condition (perfB2 > perfC2);  
  
    var perfA3 = gaussian(skillA, 15), perfC3 = gaussian(skillC, 15);  
    condition (perfA3 > perfC3);  
  
    return skillA;  
}
```

```
var res = Infer({method: 'MCMC', samples:  
                50000}, trueSkill)  
print("Expected value: "+expectation(res));  
viz.auto(res);
```

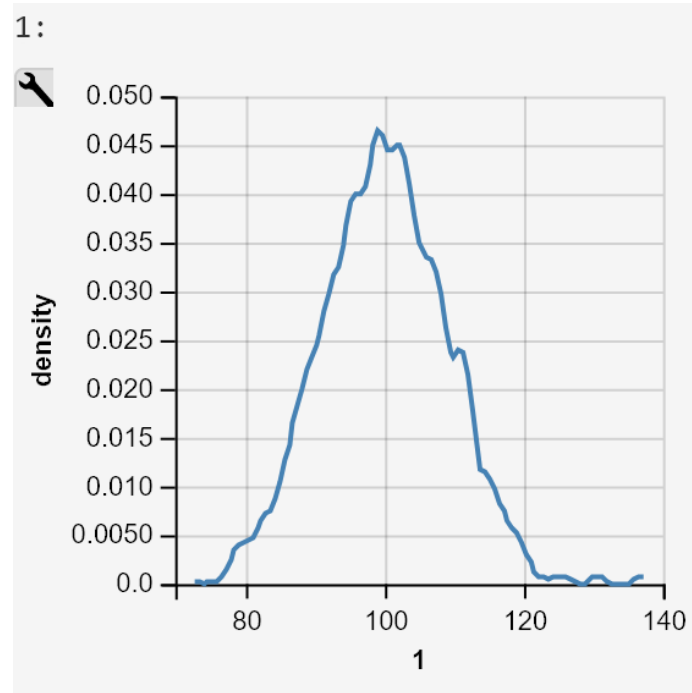
Performances

Player A



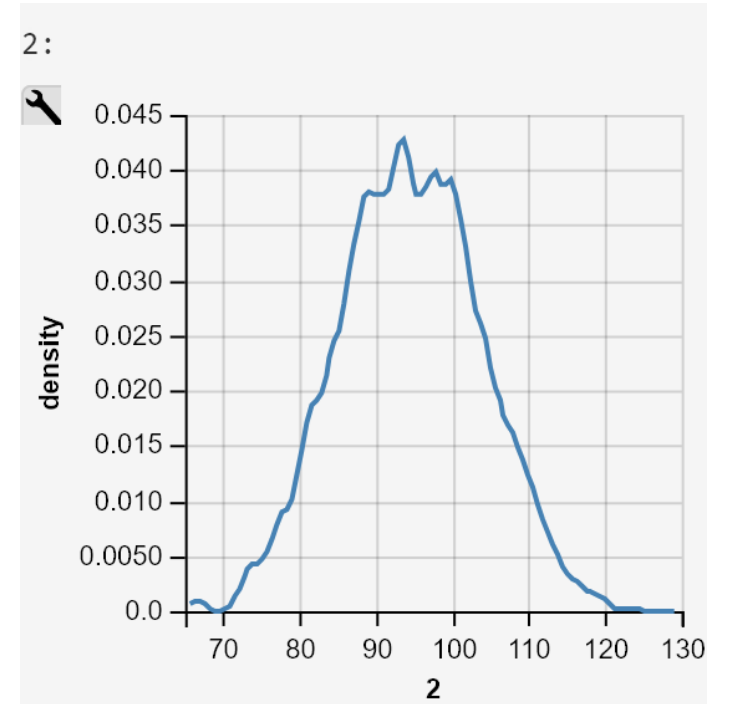
Mean: 105.3

Player B



100.4

Player C



94.6

Try at Home

Extend the problem to n games:

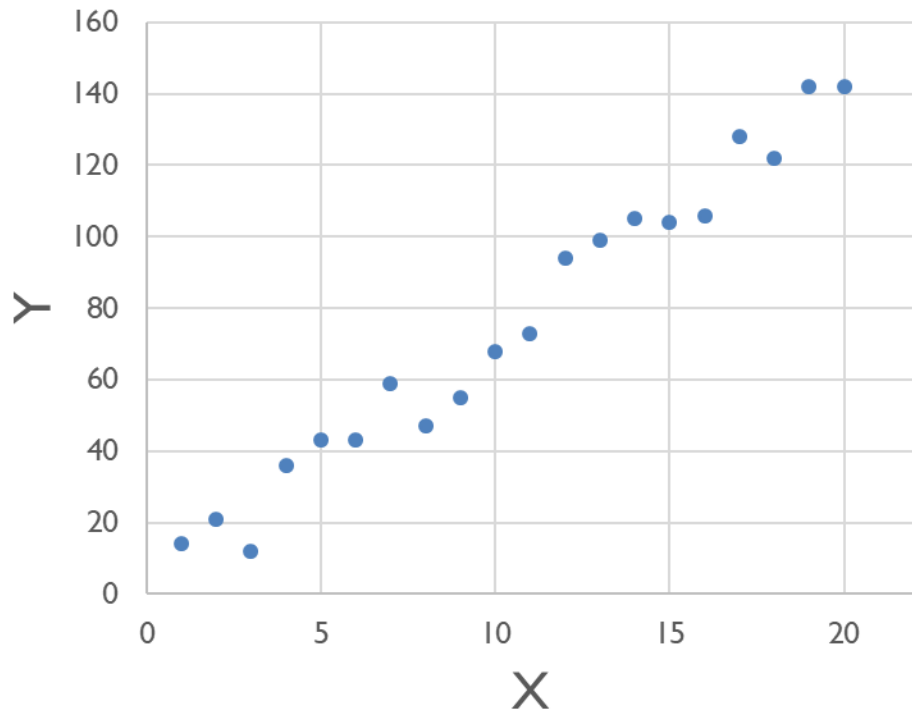
- The results array contains plays and whether 1st won e.g.,
Res = [['A', 'B', true], ['A', 'B', false], ...]
- Hint: define map M = {'A': skillA, 'B': skillB, 'C': skillC}
- Performance is now indexed by M[Res[0][0]] and M[Res[0][1]]
- Use map operator to condition on all elements of Res

Extend the problem to m players

- Use map operator to generate skills (may use number instead of letter for the player's name)

Continuous Models: Linear Regression

Given a set of points, find a linear relationship that most accurately describes this set

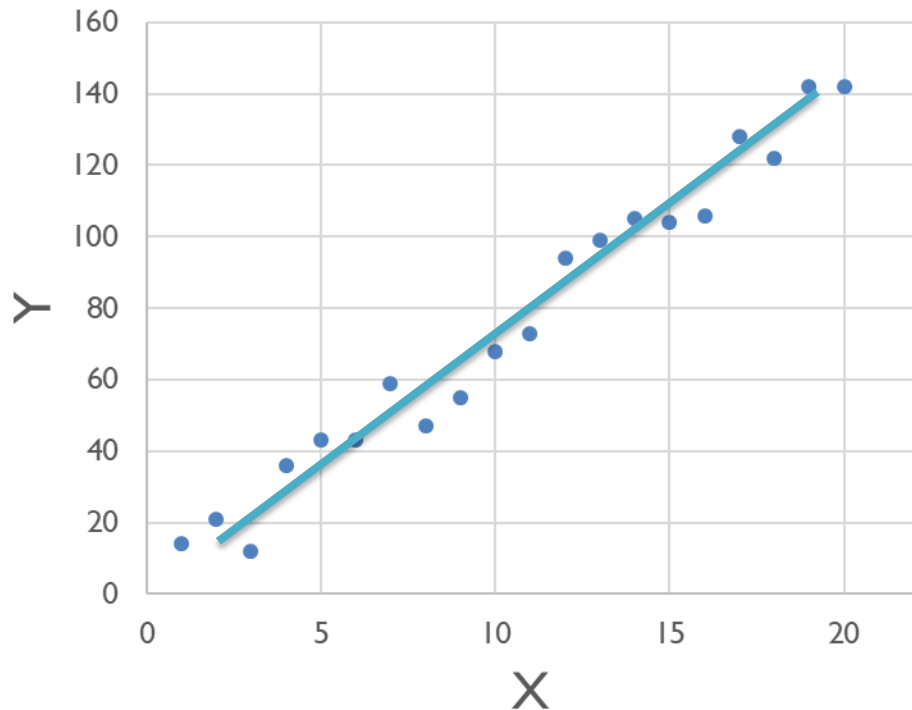


$$Y = w \cdot X + b$$

Slope \nearrow Intercept

Classical Solution (MLE)

Given a set of points, find a linear relationship that most accurately describes this set



$$Y = \mathbf{w} \cdot X + \mathbf{b}$$

Seeks point estimates (individual values) of w and b that minimize the error (e.g. square error) between the line and all the points

For the formulas and derivation see e.g.

<https://www.stat.cmu.edu/~cshalizi/mreg/15/lectures/06/lecture-06.pdf>

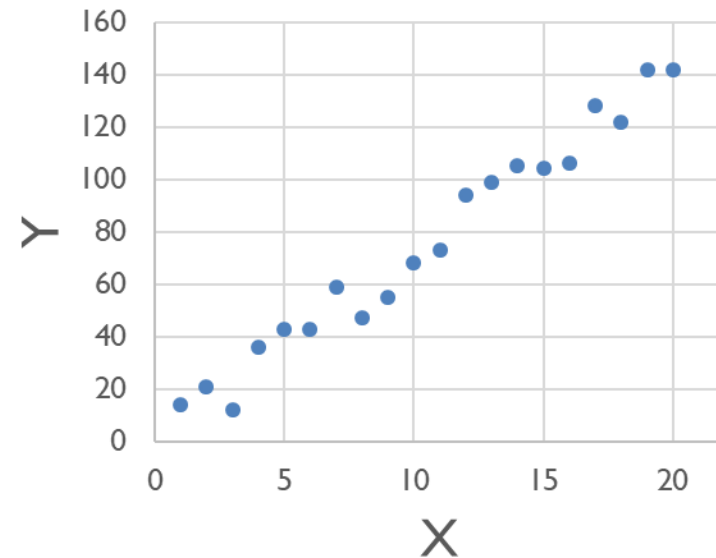
Linear Regression

```
x : [1.0, 2.0, ... ];  
y : [7.01, 14.2, .... ];
```

Datasets

```
w ~ Normal(6 , 10);  
b ~ Normal(1 , 5);  
observe(y==Normal(w*x + b, 1.0));  
posterior w;  
posterior b;
```

$$Y = w.X + b$$



Linear Regression Model

Linear Regression

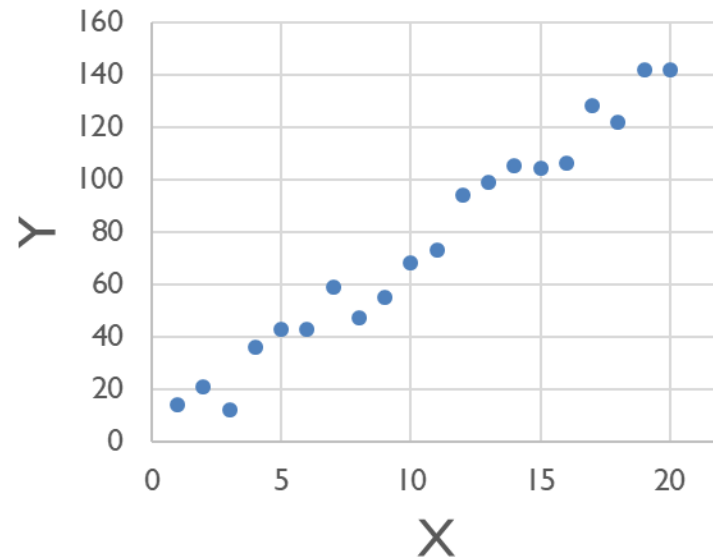
```
x : [1.0, 2.0, ... ];  
y : [7.01, 14.2, .... ];
```

Datasets

```
w ~ Normal(6 , 10);  
b ~ Normal(1 , 5);  
observe(y==Normal(w*x + b, 1.0));  
posterior w;  
posterior b;
```

Priors

$$Y = w.X + b$$



Linear Regression Model

Linear Regression

```
x : [1.0, 2.0, ... ];  
y : [7.01, 14.2, .... ];
```

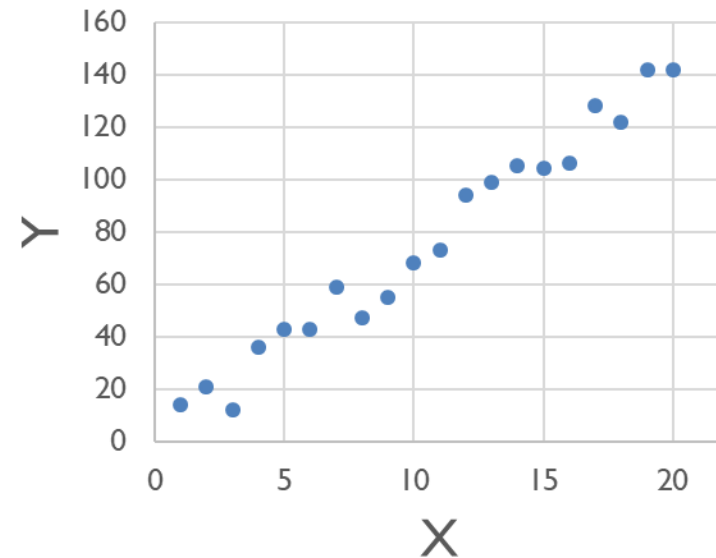
Datasets

```
w ~ Normal(6 , 10);  
b ~ Normal(1 , 5);  
observe(y==Normal(w*x + b, 1.0));  
posterior w;  
posterior b;
```

Priors

Conditioning
on Data

$$Y = w \cdot X + b$$



Linear Regression Model

Linear Regression

```
x : [1.0, 2.0, ... ];  
y : [7.01, 14.2, .... ];
```

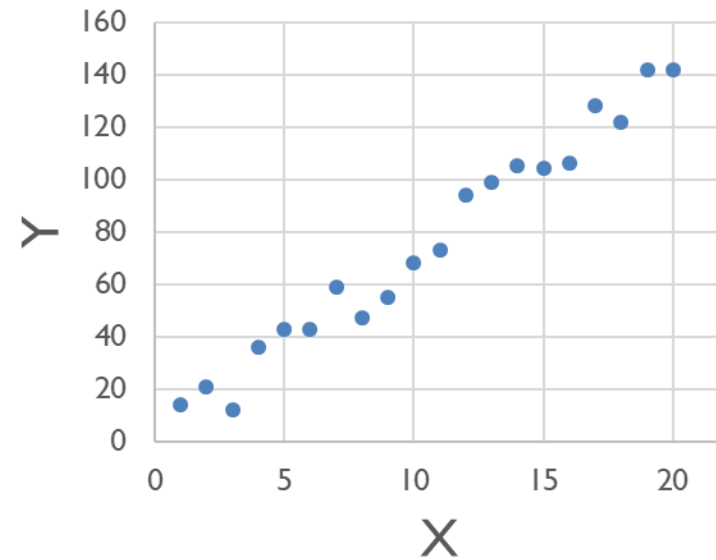
Datasets

```
w ~ Normal(6 , 10);  
b ~ Normal(1 , 5);  
observe(y==Normal(w*x + b, 1.0));  
posterior w;  
posterior b;
```

Priors

Conditioning
on Data
Queries

$$Y = w.X + b$$



Linear Regression Model

Linear Regression

```
x : [1.0, 2.0, ... ];  
y : [7.01, 14.2, .... ];
```

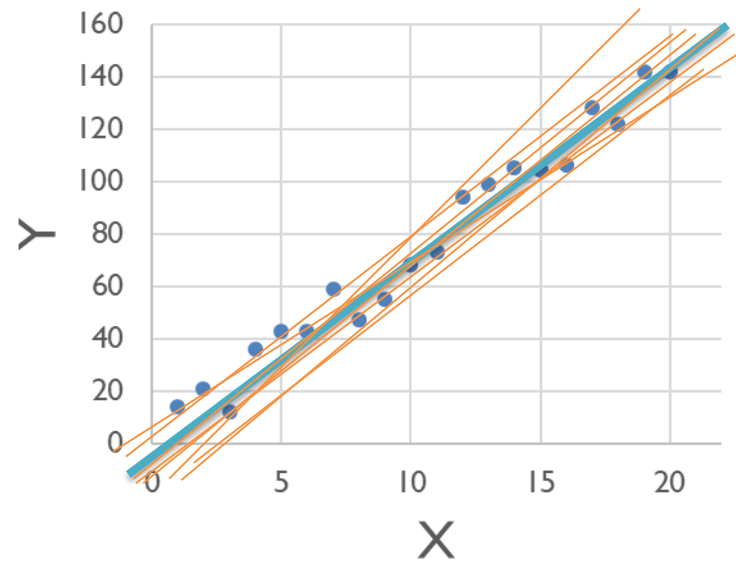
Datasets

```
w ~ Normal(6 , 10);  
b ~ Normal(1 , 5);  
observe(y==Normal(w*x + b, 1.0));  
posterior w;  
posterior b;
```

Priors

Conditioning
on Data
Queries

$$Y = w.X + b$$



Linear Regression Model

Continuous Models: Linear Regression

```
var xs = [0, 1, 2, 3]; var ys = [0, 1, 4, 6];
```

```
var model = function() {  
  var slope = gaussian(0, 2);  
  var intercept = gaussian(0, 2);  
  var sigma = 1; // for more interesting result, change to gamma(1, 1);  
  
  var f = function(x) { return slope * x + intercept; };  
  
  map2(  
    // function(x, y) { factor(Gaussian({mu: f(x), sigma: sigma}).score(y)); }, xs, ys);  
    function(x, y) { observe(Gaussian({mu: f(x), sigma: sigma}), y); }, xs, ys); //same  
  
  return [slope,intercept];  
}  
  
viz.marginals({method: 'MCMC', samples: 10000}, model));
```

Likelihood or Log-likelihood?

target = 0.0

X := Gaussian(0, 1);

target = $\mathcal{N}_{\text{logpdf}}(X, 0, 1)$

Likelihood or Log-likelihood?

target = 0.0

X := Gaussian(0, 1);

Y := Gaussian(X, 1);

target = $\mathcal{N}_{\text{logpdf}}(X, 0, 1) + \mathcal{N}_{\text{logpdf}}(Y, X, 1)$

Likelihood or Log-likelihood?

target = 0.0

X := Gaussian(0, 1);

Y := Gaussian(X, 1);

target = $\mathcal{N}_{\text{logpdf}}(X, 0, 1) + \mathcal{N}_{\text{logpdf}}(Y, X, 1)$

observe (Y, 0.5);

return X;

The probability that a continuous random variable has a particular value is equal to 0 (and $\text{logpdf}(0) = -\text{inf}$)

Likelihood or Log-likelihood?

target = 0.0

X := Gaussian(0, 1);

Y := Gaussian(X, 1);

target = $\mathcal{N}_{\text{logpdf}}(X, 0, 1) + \mathcal{N}_{\text{logpdf}}(Y, X, 1)$

observe (Y, 0.5);



factor (Gaussian.score(0.5, Y));

Likelihood or Log-likelihood?

target = 0.0

X := Gaussian(0, 1);

Y := Gaussian(X, 1);

target = $\mathcal{N}_{\text{logpdf}}(X, 0, 1) + \mathcal{N}_{\text{logpdf}}(Y, X, 1)$

factor (Gaussian.score(0.5, Y));

**target = $\mathcal{N}_{\text{logpdf}}(X, 0, 1) + \mathcal{N}_{\text{logpdf}}(Y, X, 1)$
+ $\mathcal{N}_{\text{logpdf}}(0.5, X, 1)$**

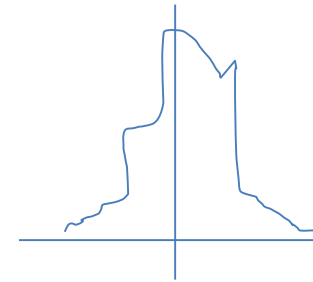
Approaches to Probabilistic Inference

Exact and Approximate

Sampling

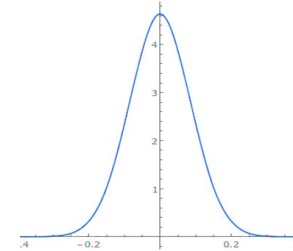
(Rejection – Church)

(MCMC – Church & Stan & R2)



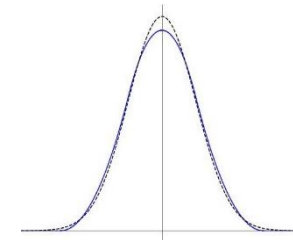
Variational Inference

(Fun & Infer.NET)

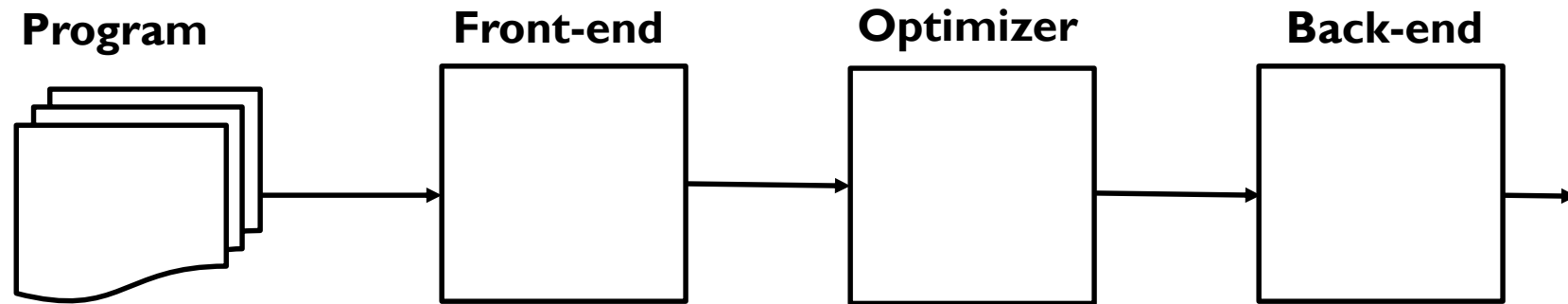


Exact Symbolic Inference

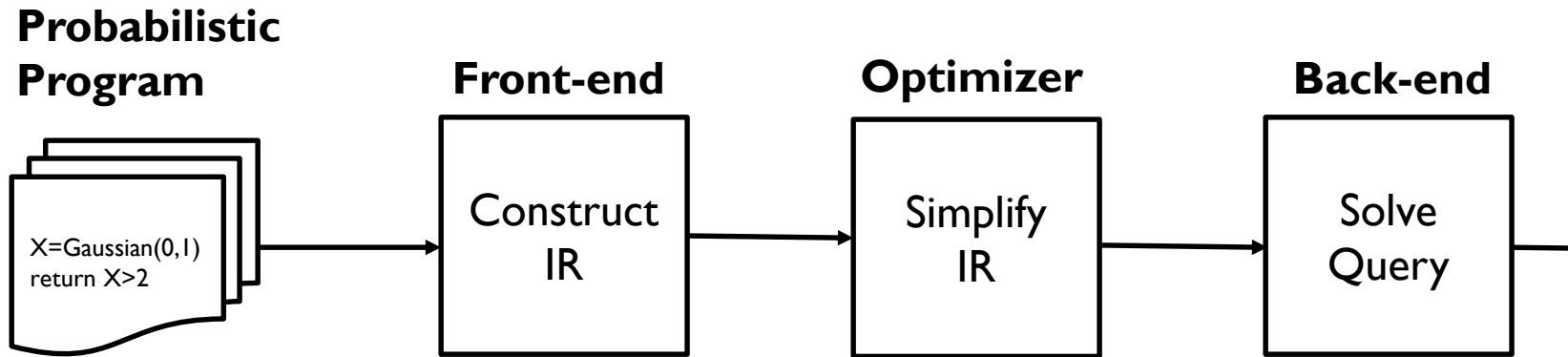
(PSI & Hakaru)



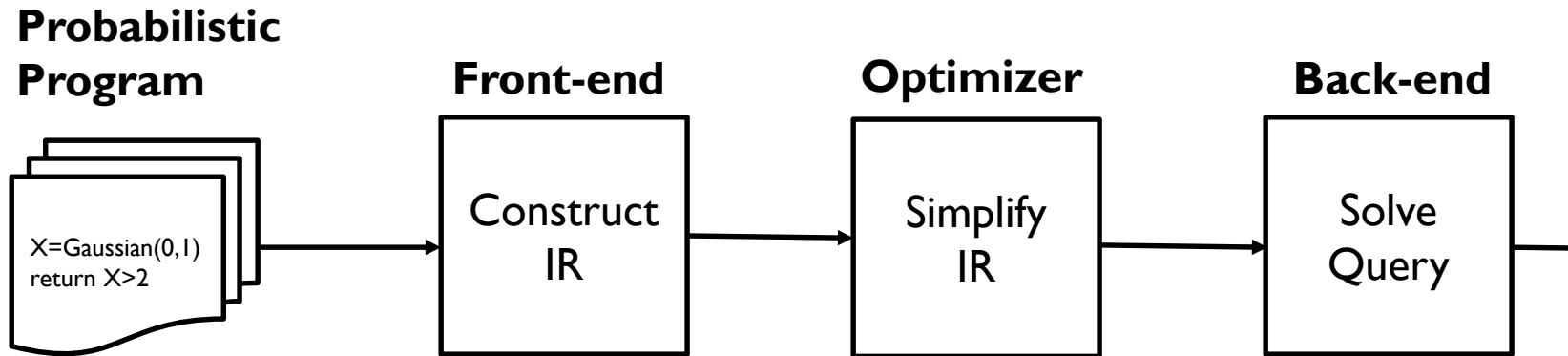
PSI Overview



PSI Overview


$$e \in E ::= x \mid e \mid \pi \mid 0 \mid 1 \mid 2 \mid \dots$$
$$\mid \log(e) \mid -e \mid e_1 + \dots + e_n \mid e_1 \cdot \dots \cdot e_n \mid e_1^{e_2}$$
$$\mid \delta(e) \mid [e_1 = e_2] \mid [e_1 \leq e_2] \mid [e_1 \neq e_2] \mid [e_1 < e_2]$$
$$\mid \int_{\mathbb{R}} dx e[x] \mid \sum_{x \in \mathbb{Z}} e[x] \mid \varphi(e_1, \dots, e_n)$$

PSI Overview



- ▶ Basic algebraic simplifications

$$x + x \rightarrow 2 \cdot x, \quad x \cdot x \rightarrow x^2, \dots$$

- ▶ Simplifications on constraints

$$[x = 0] + [x \neq 0] \rightarrow 1, \quad [x \leq 0] \cdot [0 \leq x] \rightarrow [x = 0],$$

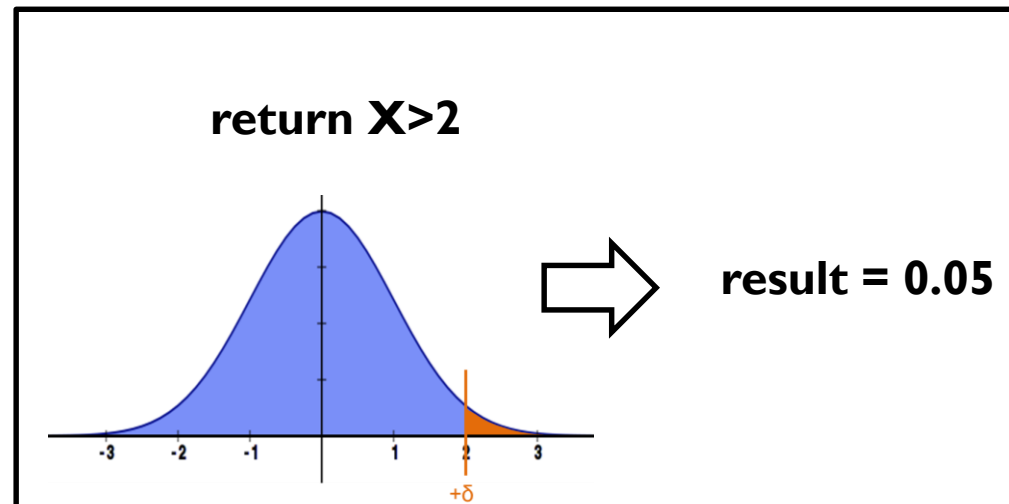
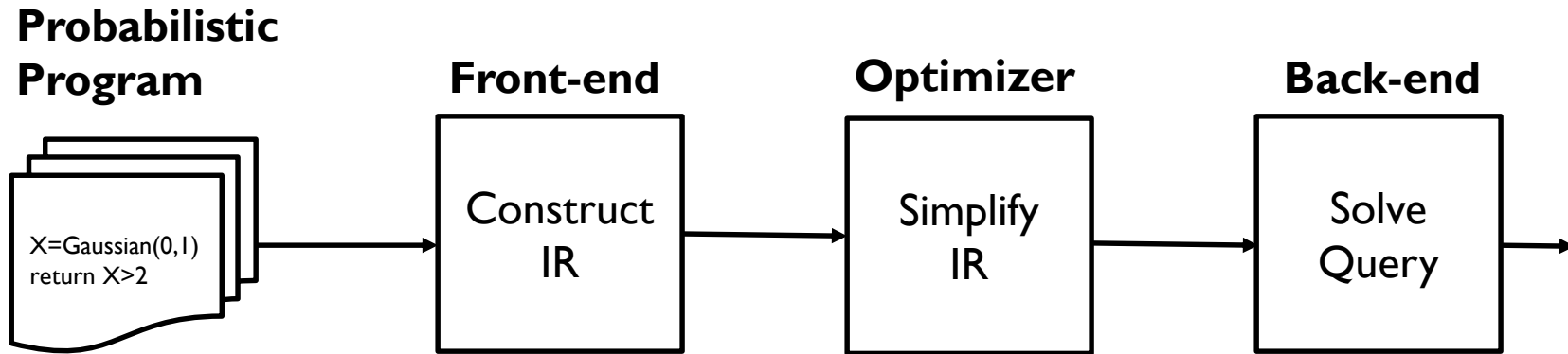
$$\delta(x) \cdot [1 \leq x] \rightarrow 0, \dots$$

- ▶ Symbolic Integration

$$\int_{\mathbb{R}} dx \int_{\mathbb{R}} dy [0 \leq x] \cdot [x \leq 1] \cdot [0 \leq y] \cdot [y \leq 1] \cdot \delta(z - x \cdot y) \rightarrow$$

$$-[0 < z] \cdot [z \leq 1] \cdot \log(z), \dots$$

PSI Overview



PSI's Symbolic Domain

$$\begin{aligned} e \in E ::= & x \mid e \mid \pi \mid 0 \mid 1 \mid 2 \mid \dots \\ & \mid \log(e) \mid -e \mid e_1 + \dots + e_n \mid e_1 \cdot \dots \cdot e_n \mid e_1^{e_2} \\ & \mid \delta(e) \mid [e_1 = e_2] \mid [e_1 \leq e_2] \mid [e_1 \neq e_2] \mid [e_1 < e_2] \end{aligned}$$

Encodes probability density functions

PSI's Symbolic Domain

$$e \in E ::= x | e | \pi | 0 | 1 | 2 | \dots$$

$$| \log(e) | -e | e_1 + \dots + e_n | e_1 \cdot \dots \cdot e_n | e_1^{e_2}$$

$$| \delta(e) | [e_1 = e_2] | [e_1 \leq e_2] | [e_1 \neq e_2] | [e_1 < e_2]$$

$$\delta(x) = \begin{cases} +\infty, & x = 0 \\ 0, & x \neq 0 \end{cases} \quad \text{and} \quad \int_{-\infty}^{\infty} \delta(x) dx = 1$$

Encodes probability density functions



Transforming Variables Using the Dirac Generalized Function

Chi AU and Judy TAM

This article provides an alternative method of finding the distribution of a function of one or more random variables using the Dirac generalized function. Unlike the conventional change-of-variable technique which involves a one-to-one transformation and computation of the Jacobian, here the procedure for obtaining the distributions of functions of random variables is shown to be simple, direct, and powerful.

KEY WORDS: Distribution of function of random variables; Change-of-variable technique; Dirac function.

In fact, this table immediately gives the probability distributions for $Y = X - 2$, and $Y = -3X$.

But for $Y = X^2$, it is given

Probability	$\frac{1}{10}$	$\frac{3}{10}$	$\frac{2}{10}$	$\frac{2}{5}$
$y = x^2$	0	1	4	16

That is to say,

$$P([X - 2 = Y = -3]) = P([X = -1]) = \frac{1}{5},$$

$$P([-3X = Y = 3]) = P([X = -1]) = \frac{1}{5},$$

and

PSI's Symbolic Domain

$$e \in E ::= x | e | \pi | 0 | 1 | 2 | \dots \\ | \log(e) | - e | e_1 + \dots + e_n | e_1 \cdot \dots \cdot e_n | e_1^{e_2} \\ | \delta(e) | [e_1 = e_2] | [e_1 \leq e_2] | [e_1 \neq e_2] | [e_1 < e_2]$$

Encodes probability density functions

$$\text{Bernoulli}(x; p) = p \cdot \delta(1 - x) + (1 - p) \cdot \delta(x)$$

PSI's Symbolic Domain

$$\begin{aligned} e \in E ::= & \quad x \mid e \mid \pi \mid 0 \mid 1 \mid 2 \mid \dots \\ & \quad \mid \log(e) \mid -e \mid e_1 + \dots + e_n \mid e_1 \cdot \dots \cdot e_n \mid e_1^{e_2} \\ & \quad \mid \delta(e) \mid [e_1 = e_2] \mid [e_1 \leq e_2] \mid [e_1 \neq e_2] \mid [e_1 < e_2] \end{aligned}$$

Encodes probability density functions

$$\text{Bernoulli}(x; p) = p \cdot \delta(1 - x) + (1 - p) \cdot \delta(x)$$

$$\text{Gauss}(x; \mu, \nu) = [\nu = 0] \cdot \delta(x - \mu) + [\nu \neq 0] \cdot \frac{e^{-(x-\mu)^2/(2\nu)}}{(2\pi\nu)^{\frac{1}{2}}}$$

PSI's Symbolic Domain

$$\begin{aligned} e \in E ::= & \quad x \mid e \mid \pi \mid 0 \mid 1 \mid 2 \mid \dots \\ & \mid \log(e) \mid -e \mid e_1 + \dots + e_n \mid e_1 \cdot \dots \cdot e_n \mid e_1^{e_2} \\ & \mid \delta(e) \mid [e_1 = e_2] \mid [e_1 \leq e_2] \mid [e_1 \neq e_2] \mid [e_1 < e_2] \\ & \mid \int_{\mathbb{R}} dx e[x] \mid \sum_{x \in \mathbb{Z}} e[x] \mid \varphi(e_1, \dots, e_n) \end{aligned}$$

Encodes probability density functions

$$\text{Bernoulli}(x; p) = p \cdot \delta(1 - x) + (1 - p) \cdot \delta(x)$$

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PSI's Symbolic Domain

$$\begin{aligned} e \in E ::= & \quad x \mid e \mid \pi \mid 0 \mid 1 \mid 2 \mid \dots \\ & \mid \log(e) \mid -e \mid e_1 + \dots + e_n \mid e_1 \cdot \dots \cdot e_n \mid e_1^{e_2} \\ & \mid \delta(e) \mid [e_1 = e_2] \mid [e_1 \leq e_2] \mid [e_1 \neq e_2] \mid [e_1 < e_2] \\ & \mid \int_{\mathbb{R}} dx e[[x]] \mid \sum_{x \in \mathbb{Z}} e[[x]] \mid \varphi(e_1, \dots, e_n) \end{aligned}$$

Encodes probability density functions

$$\text{Bernoulli}(x; p) = p \cdot \delta(1 - x) + (1 - p) \cdot \delta(x)$$

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$$\text{UniformInt}(x; a, b) = \frac{\sum_{x' \in \mathbb{Z}} \delta(x - x') \cdot [a \leq x'] \cdot [x' \leq b]}{\sum_{x' \in \mathbb{Z}} [a \leq x'] \cdot [x' \leq b]}$$

PSI's Symbolic Domain

$$\begin{aligned} e \in E ::= & \quad x \mid e \mid \pi \mid 0 \mid 1 \mid 2 \mid \dots \\ & \mid \log(e) \mid -e \mid e_1 + \dots + e_n \mid e_1 \cdot \dots \cdot e_n \mid e_1^{e_2} \\ & \mid \delta(e) \mid [e_1 = e_2] \mid [e_1 \leq e_2] \mid [e_1 \neq e_2] \mid [e_1 < e_2] \\ & \mid \int_{\mathbb{R}} dx e[x] \mid \sum_{x \in \mathbb{Z}} e[x] \mid \varphi(e_1, \dots, e_n) \\ & \mid (d/dx)^{-1}[e^{-x^2}](e) \quad (\text{Error function}) \end{aligned}$$

Encodes probability density functions

$$\text{Bernoulli}(x; p) = p \cdot \delta(1 - x) + (1 - p) \cdot \delta(x)$$

$$\text{Gauss}(x; \mu, \nu) = [\nu = 0] \cdot \delta(x - \mu) + [\nu \neq 0] \cdot \frac{e^{-(x-\mu)^2/(2\nu)}}{(2\pi\nu)^{\frac{1}{2}}}$$

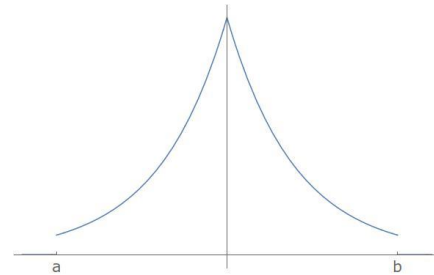
$$\text{UniformInt}(x; a, b) = \frac{\sum_{x' \in \mathbb{Z}} \delta(x - x') \cdot [a \leq x'] \cdot [x' \leq b]}{\sum_{x' \in \mathbb{Z}} [a \leq x'] \cdot [x' \leq b]}$$

How About Sum of **Three** Variables?

$$Z = X_1 + X_2 + X_3$$

X_1, X_2, X_3
i.i.d.

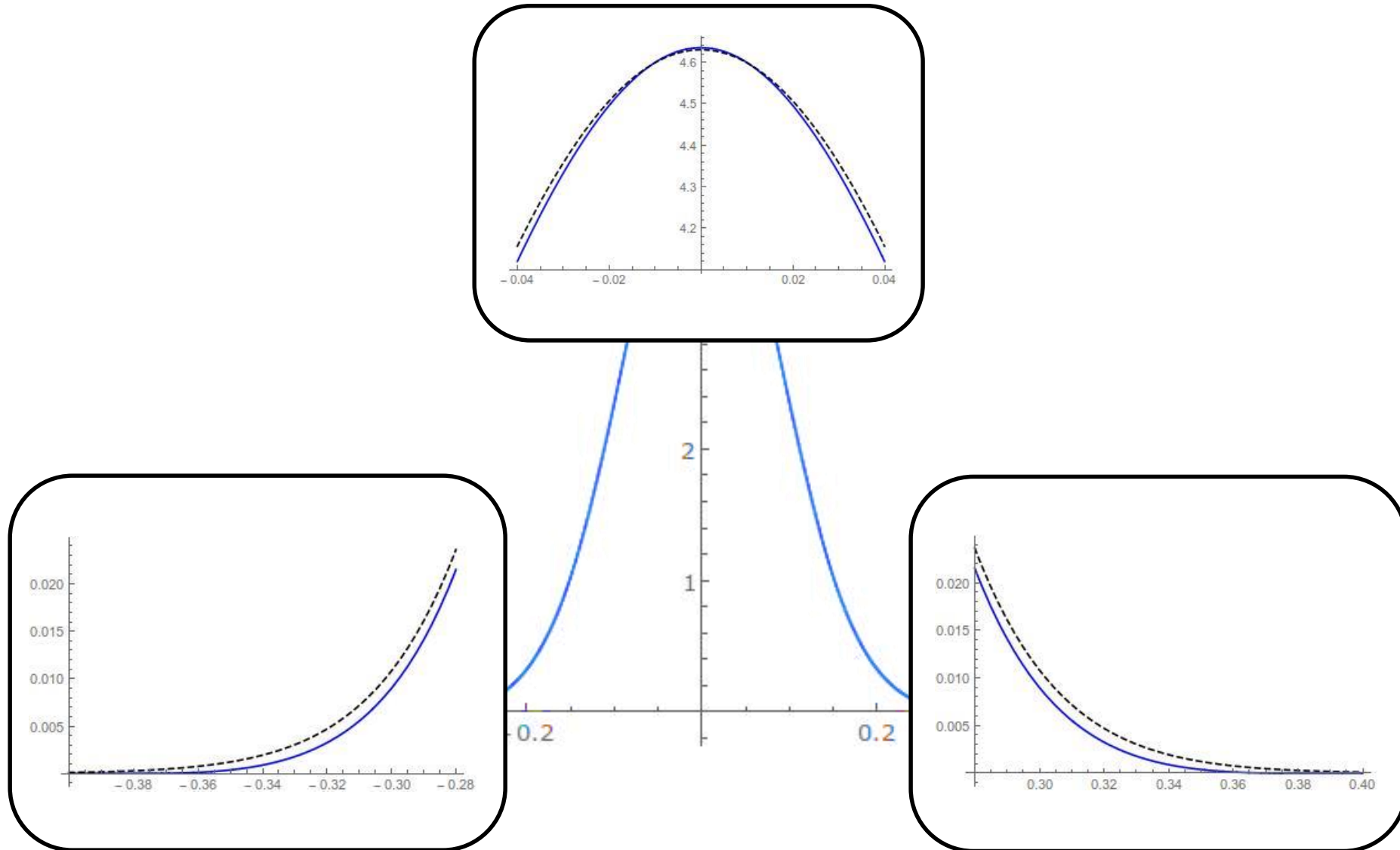
Truncated Laplace



Z



Exact Final Distribution (PSI)

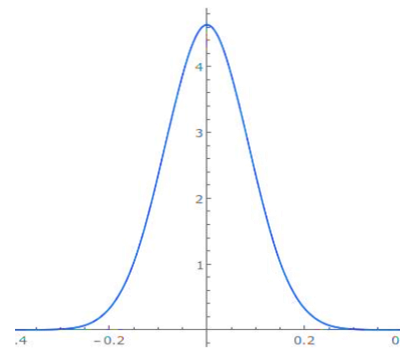
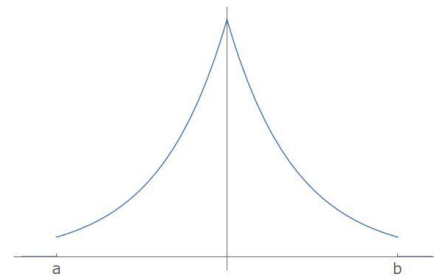


Different Input Assumptions

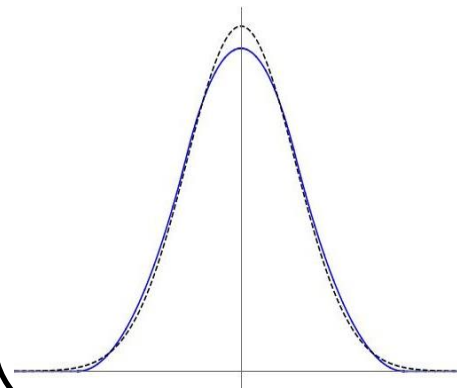
X_1, X_2, X_3
i.i.d.

Z

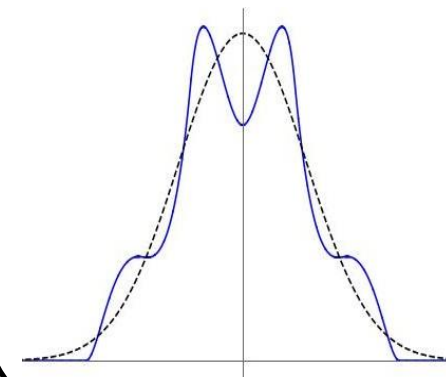
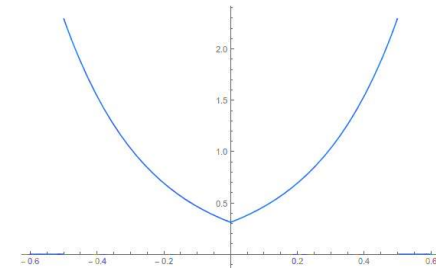
Truncated Laplace



Uniform



Reverse Exponential



Continuous Example

$X := \text{Laplace}(0, 1)$

$Y := \text{Laplace}(0, 1)$

$Z = X + Y$

observe $Z > 2$

return X

Continuous Example

X := Laplace(0,1)

Y := Laplace(0,1)

Z = X + Y

observe Z > 2

return X

Variable Definition Rule:

0: Add variable to the state:

$$p() \rightarrow p(X)$$

1: Get the Laplace distribution expression

$$f(t_{tmp}) = \frac{\lambda}{2} e^{-\lambda|t_{tmp}-\mu|}$$

2: Relate it to the program's variable X

$$p(X) = \text{replace } t_{tmp} \text{ with } X \text{ in } f(t_{tmp})$$

Continuous Example

X := Laplace(0,1)

Y := Laplace(0,1)

Z = X + Y

observe Z > 2

return X

Variable Definition Rule:

0: Add variable to the state:

$$p() \rightarrow p(X)$$

1: Get the Laplace distribution expression

$$f(t_{tmp}) = \frac{\lambda}{2} e^{-\lambda|t_{tmp}-\mu|}$$

2: Relate it to the program's variable X

$$p(X) = \int dt_{tmp} \cdot \delta[X - t_{tmp}] \cdot f(t_{tmp})$$

Continuous Example

X := Laplace(0,1)

$$p(X) = \int dt_{tmp} \cdot \delta[X - t_{tmp}] \cdot \frac{\lambda}{2} e^{-\lambda|t_{tmp}-\mu|}$$

Y := Laplace(0,1)

Z = X + Y

observe Z > 2

return X

Integration with Dirac Deltas:

$$\int dt_{tmp} \cdot \delta[X - t_{tmp}] \cdot f(t_{tmp}) = f(X)$$

Continuous Example

X := Laplace(0, 1)

$$p(X) = \frac{\lambda}{2} e^{-\lambda|X-\mu|}$$

Y := Laplace(0, 1)

Z = X + Y

observe Z > 2

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Integration with Dirac Deltas:

$$\int dt_{tmp} \cdot \delta[X - t_{tmp}] \cdot f(t_{tmp}) = f(X)$$

Continuous Example

X := Laplace(0,1)

$$p(X) = \frac{1}{2} e^{-|X|}$$

Y := Laplace(0,1)

Z = X + Y

observe Z > 2

return X

Integration with Dirac Deltas:

$$\int dt_{tmp} \cdot \delta[X - t_{tmp}] \cdot f(t_{tmp}) = f(X)$$

Continuous Example

X := Laplace(0,1)

$$p(X) = \frac{1}{2} e^{-|X|}$$

Y := Laplace(0,1)

$$p(X, Y) = \frac{1}{2} e^{-|X|} \cdot \frac{1}{2} e^{-|Y|}$$

Z = X + Y

observe Z > 2

return X

Continuous Example

X := Laplace(0, 1)

$$p(X) = \frac{1}{2} e^{-|X|}$$

Y := Laplace(0, 1)

$$p(X, Y) = \frac{1}{2} e^{-|X|} \cdot \frac{1}{2} e^{-|Y|}$$

Z = X + Y

$$p(X, Y, Z) = \frac{1}{2} e^{-|X|} \cdot \frac{1}{2} e^{-|Y|} \cdot \delta(Z - (X + Y))$$

observe Z > 2

return X

Assignment Rule (Deterministic):

0: Add variable Z to the state:

$$p(X, Y) \rightarrow p(X, Y, Z)$$

I: Relate the program's variables

$$p(X, Y, Z) = p(X, Y) \cdot \delta[Z - Expr(X, Y)]$$

Continuous Example

X := Laplace(0, 1)

$$p(X) = \frac{1}{2} e^{-|X|}$$

Y := Laplace(0, 1)

$$p(X, Y) = \frac{1}{2} e^{-|X|} \cdot \frac{1}{2} e^{-|Y|}$$

Z = X + Y

$$p(X, Y, Z) = \frac{1}{2} e^{-|X|} \cdot \frac{1}{2} e^{-|Y|} \cdot \delta(Z - (X + Y))$$

observe Z > 2

$$p(X, Y, Z | Z > 2) \sim \frac{1}{2} e^{-|X|} \cdot \frac{1}{2} e^{-|Y|} \cdot \delta(Z - (X + Y)) \cdot [Z > 2]$$

return X

Continuous Example

X := Laplace(0, 1)

$$p(X) = \frac{1}{2} e^{-|X|}$$

Y := Laplace(0, 1)

$$p(X, Y) = \frac{1}{2} e^{-|X|} \cdot \frac{1}{2} e^{-|Y|}$$

Z = X + Y

$$p(X, Y, Z) = \frac{1}{2} e^{-|X|} \cdot \frac{1}{2} e^{-|Y|} \cdot \delta(Z - (X + Y))$$

observe **Z** > 2

$$p(X, Y, Z | Z > 2) \sim \frac{1}{2} e^{-|X|} \cdot \frac{1}{2} e^{-|Y|} \cdot \delta(Z - (X + Y)) \cdot [Z > 2]$$

return **X**



1. Marginalization

2. Normalization

$$p(X | Z > 2) \sim \int dY \int dZ p(X, Y, Z | Z > 2)$$

$$Const = \int dX \int dY \int dZ p(X, Y, Z | Z > 2)$$

Continuous Example

I. Marginalization

$$\begin{aligned} p(X | Z > 2) &\sim \int dY \int dZ p(X, Y, Z | Z > 2) \\ &\sim \int dY \int dZ \frac{1}{2} e^{-|X|} \cdot \frac{1}{2} e^{-|Y|} \cdot \delta(Z - (X + Y)) \cdot [Z > 2] \\ &\sim \int dY \frac{1}{2} e^{-|X|} \cdot \frac{1}{2} e^{-|Y|} \cdot [X + Y > 2] \\ &\sim \frac{1}{4} e^{-|X|} \cdot (e^{X-2} \cdot [X \leq 2] + (2 - e^{2-X}) \cdot [X > 2]) \end{aligned}$$

2. Normalization

$$Const = \int dx \int dy \int dz p(X, Y, Z | Z > 2) = e^{-2}$$

Simplification Rules



- Unpredictable
- Sometimes slow
- Sometimes wrong

PSI uses its **own** symbolic reasoning

Simplification Rules

- **Basic Algebraic Simplifications**

$$x + x \rightarrow 2 \cdot x, \quad x \cdot x \rightarrow x^2, \dots$$

- **Simplifications of Constraints**

$$[x = 0] + [x \neq 0] \rightarrow 1, \quad [x \leq 0] \cdot [0 \leq x] \rightarrow [x = 0],$$

$$\delta(x) \cdot [1 \leq x] \rightarrow 0, \dots$$

- **Guard Linearization**

$$\delta(y - x^2) \rightarrow [-y \leq 0] \cdot ([x = 0] \cdot \delta(y) + [x \neq 0] \cdot \frac{1}{2\sqrt{y}} (\delta(x - \sqrt{y}) + \delta(x + \sqrt{y})))$$

- **Symbolic Integration**

$$\int_{\mathbb{R}} dx \int_{\mathbb{R}} dy [0 \leq x] \cdot [x \leq 1] \cdot [0 \leq y] \cdot [y \leq 1] \cdot \delta(z - x \cdot y) \rightarrow$$

$$-[0 < z] \cdot [z \leq 1] \cdot \log(z), \dots$$

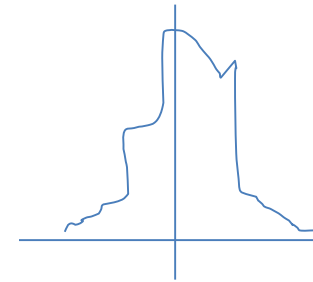
Approaches to Probabilistic Inference

Exact and Approximate

Sampling

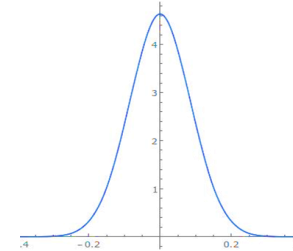
(Rejection – Church)

(MCMC – Church & Stan & R2)



Variational Inference

(Fun & Infer.NET)



Exact Symbolic Inference

(PSI & Hakaru)

